### Linear Regression Models P8111

Lecture 11

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### Today's Lecture

- Review of tests
- The bootstrap
- Permutation testing
- Cross validation

#### Individual coefficients

#### For individual coefficients

■ We can use the test statistic

$$T = \frac{\hat{\beta}_j - \beta_j}{\widehat{se}(\hat{\beta}_j)} = \frac{\hat{\beta}_j - \beta_j}{\sqrt{\hat{\sigma}^2 (X^T X)_{jj}^{-1}}} \sim t_{n-p-1}$$

• For a two-sided test of size  $\alpha$ , we reject if

$$|T| > t_{1-\alpha/2,n-p-1}$$

■ The p-value gives  $P(|t_{n-p-1}| > |T_{obs}||H_0)$ 

Note that t is a symmetric distribution that converges to a Normal as n - p - 1 increases.

#### Inference for linear combinations

Sometimes we are interested in making claims about  $c^T \beta$  for some c.

- Define  $H_0: c^T \beta = c^T \beta_0$  or  $H_0: c^T \beta = 0$
- We can use the test statistic

$$T = \frac{c^T \hat{\boldsymbol{\beta}} - c^T \boldsymbol{\beta}}{\widehat{se}(c^T \hat{\boldsymbol{\beta}})} = \frac{c^T \hat{\boldsymbol{\beta}} - c^T \boldsymbol{\beta}}{\sqrt{\hat{\sigma}^2 c^T (\boldsymbol{X}^T \boldsymbol{X})^{-1} c}}$$

- This test statistic is asymptotically Normally distributed
- For a two-sided test of size  $\alpha$ , we reject if

$$|T| > z_{1-\alpha/2}$$

#### Global F tests

■ Compute the test statistic

$$F_{obs} = \frac{(RSS_S - RSS_L)/(df_S - df_L)}{RSS_L/df_L}$$

- If  $H_0$  (the null model) is true, then  $F_{obs} \sim F_{df_S df_L, df_L}$
- Note  $df_S = n p_S 1$  and  $df_L = n p_L 1$
- We reject the null hypothesis if the p-value is above  $\alpha$ , where

$$p
-value = P(F_{df_S - df_L, df_L} > F_{obs})$$

#### The Wald test

For a vector of coefficients, we can test  $H_0: \beta = \beta_0$ :

■ Use the test statistic

$$W = (\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_0)^T [Var(\hat{\boldsymbol{\beta}})]^{-1} (\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}_0)$$

- Under the null, this test statistic has an asymptotic  $\chi_p^2$  distribution
- In practice, we replace  $Var(\hat{\beta})$  with  $\widehat{Var}(\hat{\beta})$  and use an F distribution

#### The LRT

If we are using maximum likelihood estimation (we'll cover this soon – turns out to be least squares in MLR), we can use a LRT:

Use the test statistics

$$\Delta = -2\log\frac{L_0}{L_1} = -2(l_0 - l_1)$$

■ This test statistic has an asymptotic  $\chi_d^2$  distribution where d is the difference in the number of parameters between the two models.

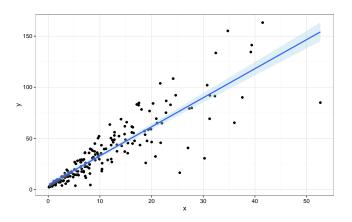
# Inference: departure from assumptions

- In large samples,  $\hat{\beta}$  is approximately Normal even if the errors are not
- In smaller samples, especially when our assumptions are not justified, the inferential methods we've developed are not valid
- Might also want variance estimates for quantities that are difficult to derive analytically

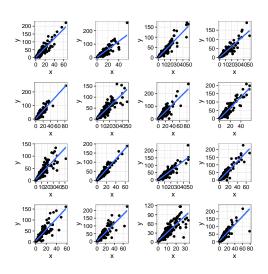
### Resampling methods

- Fundamental tool in modern statistics
- Build on the intuition that, if we could, we'd keep gathering new datasets – but we can't
- Use repeated samples of a training set to understand variability
- Computationally intensive ... but we have computers

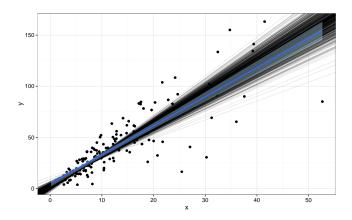
# Motivating the Bootstrap



### Motivating the Bootstrap



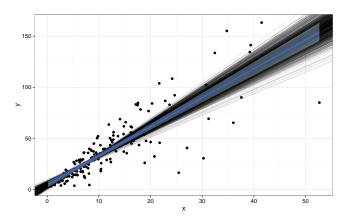
# Motivating the Bootstrap



### The Bootstrap

- The basic idea is that the observed data mimics the underlying distribution, whatever that may be
- Drawing samples (with replacement) from the observed data mimics drawing samples from the underlying distribution
- Recalculating regression parameters for the "new" samples gives an idea of the distribution of regression coefficients

# Implementing the Bootstrap

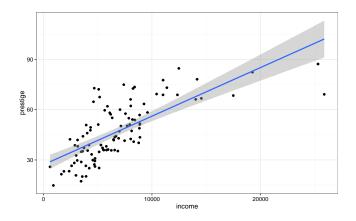


### Bootstrap example

#### Prestige dataset

- Information on 102 occupations
- Variables include education, income, proportion women, job type, and prestige
- Source: 1971 Canadian census

#### Non-normal inference

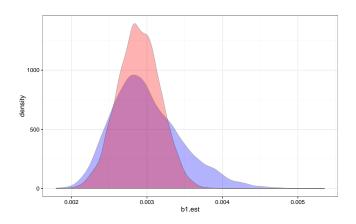


### Bootstrap code

```
## define a vector for the bootstrapped estimates
betaHatBS = data.frame(b1.est = rep(NA, 10000))

## use a loop to do the bootstrap
for(i in 1:10000) {
    data.cur = sample_frac(Prestige, size = 1, replace = TRUE)
    betaHatBS$b1.est[i] = lm(prestige ~ income, data = data.cur)$coef[2]
}
```

# Bootstrap results



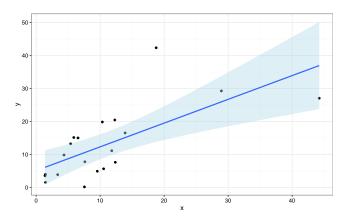
# Permutation testing

- Bootstrapping helps understand variability
- What about testing?
- One option invert the CI from a bootstrap
- Another option understand distribution of "test statistic" under the null

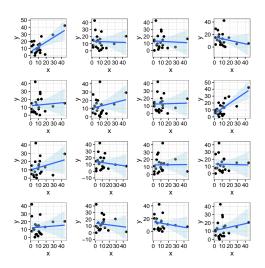
#### Permute the data

- If we permute the data, there should be no association
- Easy for comparing two groups or SLR; harder for MLRs

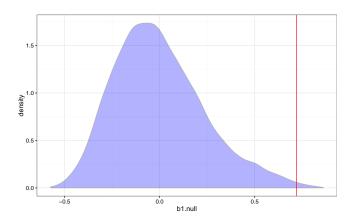
# Permutation test example



### Permutation test example



# Permutation test example



### Implementing permutation tests

```
## do enough permutations to test
obs.coef = coef(lm(y ~ x, data = data.noncst))[2]
b1 = data.frame(b1.null = rep(NA, 10000))
for(i in 1:10000){
   data.noncst.cur = mutate(data.noncst, x = sample(x, length(x), replace = FALSE))
   b1$b1.null[i] = coef(lm(y ~ x, data = data.noncst.cur))[2]
}
```

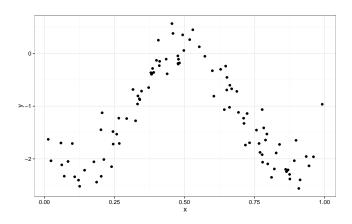
#### Cross Validation

- Focus is on model performance, quantified by prediction error
- We get in-sample performance ...
- But we want generalization to new data
- Most of the time, we don't have an external testing dataset

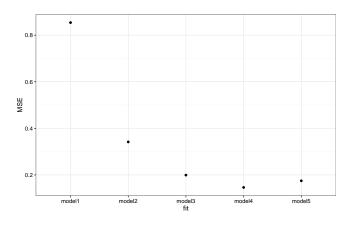
#### Cross Validation: one validation set

- Simplest case: create a validation set by randomly splitting the full dataset
- Fit model to training data; compute mean squared prediction error on test set
- Provides way of comparing models (any models ...)

# Example data

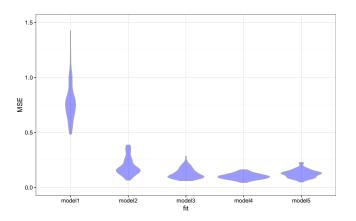


#### Cross Validation: one validation set



### Cross Validation: many validation set

 How you split the data is random; can repeat to understand this source of uncertainty



### Implementing CV

```
MSEs = data.frame(
  model1 = rep(NA, 100),
for(i in 1:100){
  set.seed(i)
  data.nonlin = mutate(data.nonlin,
                       cv_group = sample(1:100, 100, replace = FALSE) <= 80,
                       cv_group = factor(cv_group, levels = c(TRUE, FALSE),
                                         labels = c("train", "test")))
  data.train = filter(data.nonlin, cv group == "train")
  data.test = filter(data.nonlin, cv group == "test")
  fit.1 = lm(v \sim x, data = data.train)
  MSEs[i,1] = mean((data.test$y - predict(fit.1, newdata = data.test))^2)
  fit.2 = lm(y ~ x + spline_5, data = data.train)
  MSEs[i,2] = mean((data.test$y - predict(fit.2, newdata = data.test))^2)
```

#### Cross Validation: folds

- Could use *k*-fold cross validation:
  - ► Divide data into *k* equal-sized folds
  - Use each one in turn as the validation set; average MSE across sets
  - ► *k* of 5 or 10 is pretty common

# Today's big ideas

Resampling methods

■ Suggested reading: ISLR chapter 5